

Estimating the distribution of fitness effects in a structured population of cells



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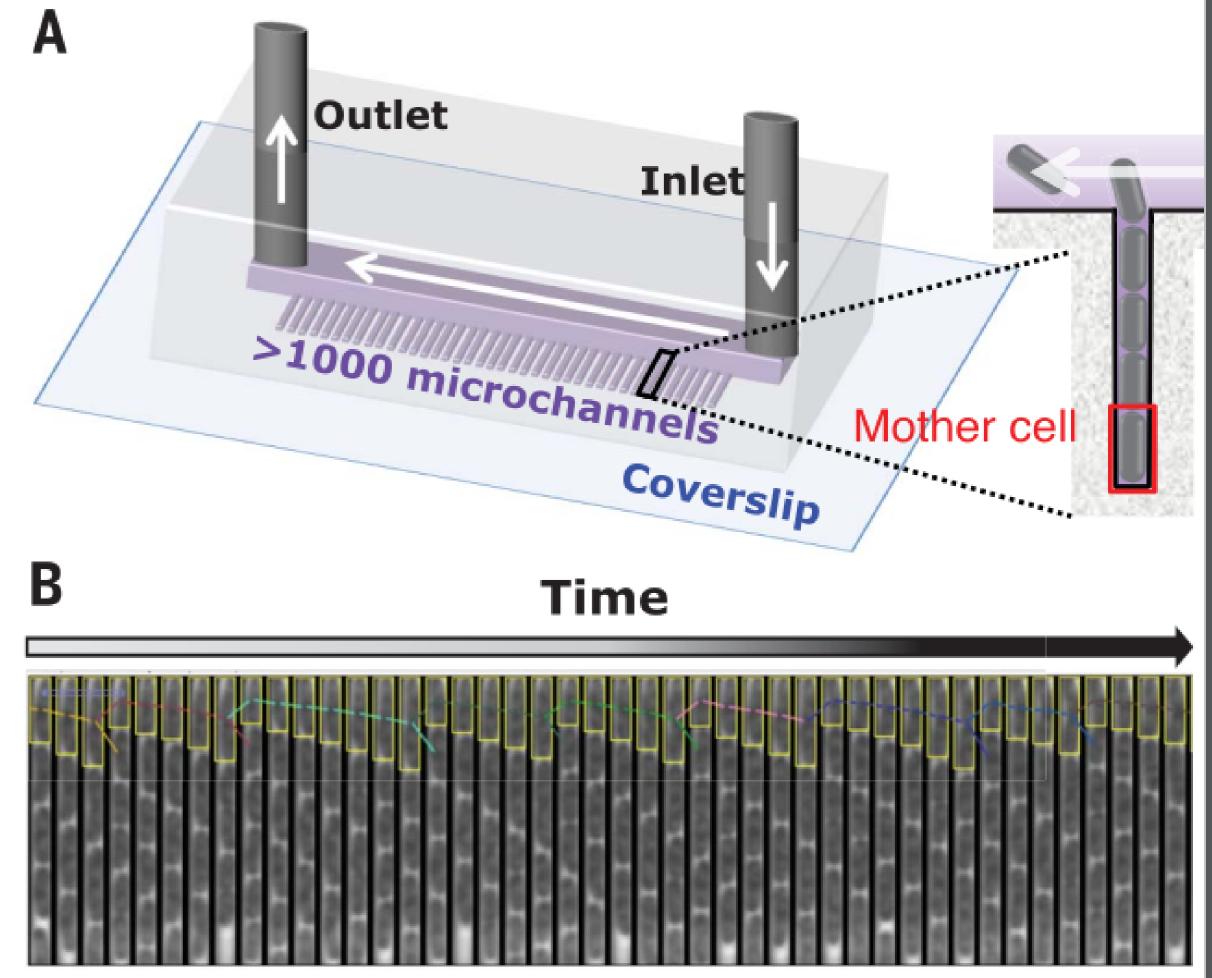
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Introduction and **Data**

- All organisms are subject to mutations
- These new traits can change the selective value (fitness) of an individual We call *Fitness* the ability of an individual with a certain genome to survive and reproduce
- How these mutations affect the selective value is a central question in evolutionary biology
- The density of the distribution of these effects is called the Distribution of Fitness Effect (DFE)

Probabilistic Model :

- 1. Z_t^J represents the noisy measure of the fitness of the cell in channel $j \in J$ at time t.
- 2. N_t^j represents the number of times the cell in channel j has mutated.



 $(N_j(t), j \ge 1)$ are *i.i.d* Poisson processes with intensity $\lambda \in (0, \infty)$.

- 3. X_k^j represents the effect of the k-th mutation on the cell in channel j. $(X_i^j)_{i,j\geq 0}$ are *i.i.d* with density $f \in L^1(\mathbb{R}) \cap L^2(\mathbb{R})$.
- 4. ε_t^j represents the measurement noise at time t for channel j. $(\varepsilon_t^j)_{j\geq 0}$ are *i.i.d* and that $\mathbb{E}(\varepsilon_t^j) = 0$.
- ► We consider a noisy compound Poisson process:

$$Z_t^j = \left(\sum_{k=1}^{N_t^j} X_k^j\right) + \varepsilon_t^j, \ t \ge 0.$$

Statement of the problem: Estimate the density of X_i from observations of Z_t on each channel $j \in J$

Statistical Strategy and statistical Results

Strategy: Estimate the characteristic function of X :

 (heuristic) If φ_X(ξ) ≃ φ̂_X(ξ), then f(x) ≃ f̂(x)

 Indeed, the characteristic function φ_X → Density f of X:

 f(x) = 1/2 ∫ φ_X(ξ)e^{-ixξ}dξ

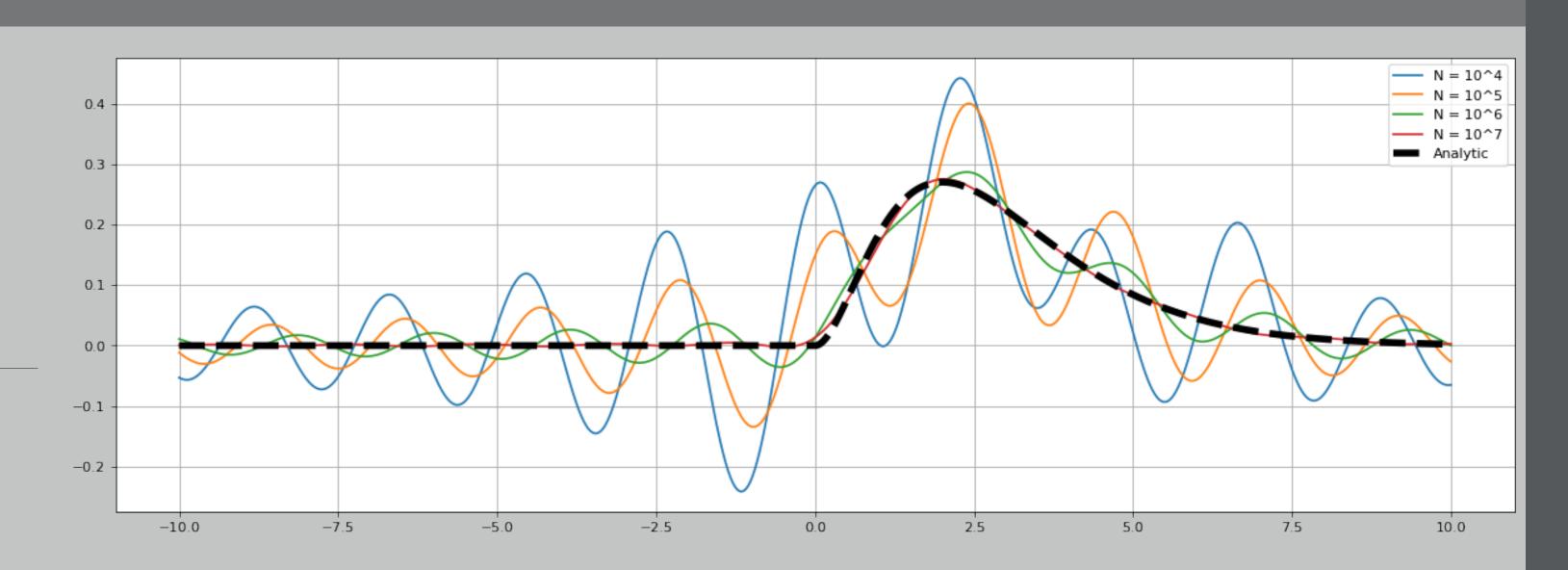


Figure 1: Measurement of the evolution of the fitness of several cell lines over time

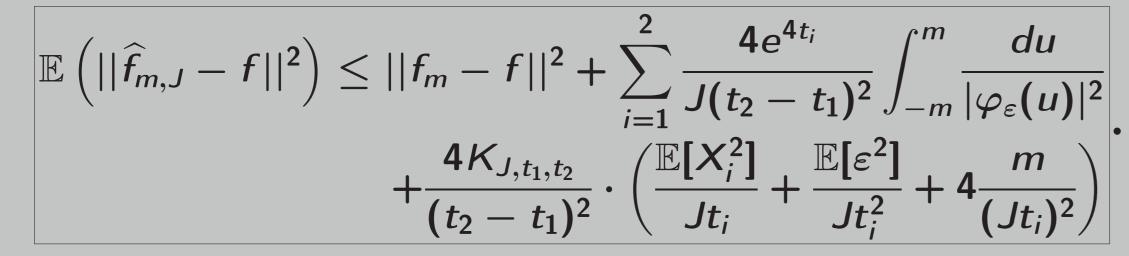
Robert et al., 2018

Combine two classical problems in non-parametric inference.

- Deconvolution
- Decompounding

 $2\pi J_{\mathbb{R}}$

▶ Theorem : For all reals $0 < t_1 < t_2$ such that $t_2 \leq \frac{1}{4} \log(Jt_2)$ $Jt_1 \rightarrow \infty, Jt_2 \rightarrow \infty$ as $J \rightarrow \infty$ and for any $m < C_{t_1,t_2}^J$, the following inequality holds



where K_{J,t_1,t_2} and C_{t_1,t_2}^J depends on m, t_1, t_2 and $\log \varphi_{\varepsilon}(\cdot)$. Futhermore, m can be chosen in an optimal way from data

A link with PDEs

The evolution of the stochastic process can also be described with PDEs.

Step 1a: If every mutation is deleterious

$$\partial u(t,x) = -\lambda u(t,x) + \int_{x}^{\infty} k\left(\frac{x}{y}\right) u(t,y) \frac{dy}{y}.$$

 $\partial u(t,x) = -\lambda u(t,x) + \int_0^\infty k\left(\frac{x}{v}\right) u(t,y)\frac{dy}{v}.$

Figure 2:Reconstruction of the Gamma $\Gamma(3)$ distribution with J channels, corrupted by a Gaussian noise $\mathcal{J}(0,1)$ with $J \in 10^4, 10^5, 10^6, 10^7$ and $t_1 = 0.1, t_2 = 1, m = 3$

The estimator converges to f when $J \rightarrow \infty$.

ex. If $k_0(dx) = (1 - \alpha)\delta_{1-\varepsilon} + \alpha\partial_{1+\eta}$, then $K(s; \alpha, \varepsilon, \eta) = (1 - \alpha)(1 - \varepsilon)^{s-1} + \alpha(1 + \eta)^{s-1}$

Step 2. Add a terme of growth ?

$$\partial u(t,x) = -\lambda u(t,x) + \int_0^\infty k\left(\frac{x}{y}\right) u(t,y)\frac{dy}{y} + xu(t,x).$$

Step 3. How to modify Step 2 to have

Where K(1) = 1 when $K(s) = \int_0^\infty x^{s-1} k(x) dx$ • Step 1b: If mutations can have some beneficial effects

Where K(1) = 1. Here, K can be non-monotonous.

$$u(t,x) \rightarrow U(x)e^{\alpha t}$$

Why do that ?

Our goal is to understand the diversity of mutation in a stable population.

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